

1. We have studied cross-sectional data so far. If we draw a new sample from the population of interest over time, then we have a pooled cross-section dataset (when new units are sampled each period) and a panel or longitudinal dataset (when we track the same units). This means we have two dimensions (units  $i$  and time periods  $t$ ) to consider our research question relating<sup>1</sup> outcome  $Y_{it}$  to treatment  $X_{it}$  possibly with access to instrument  $Z_{it}$ . Discuss the assumptions under which the Random Effects (RE) vs Fixed Effects (FE) estimators consistently estimate the true effect of treatment on the outcome:
  - Random Effects: OLS on  $Y_{it} = \theta_t + \beta \cdot X_{it} + U_{it}$  with dummy variables for time periods (possibly interacting with treatment to assess structural change over time)
  - Fixed Effects: OLS on  $Y_{it} = \alpha_i + \theta_t + \beta \cdot X_{it} + U_{it}$  with dummy variables for time periods and individual units (if we have a panel) or exogenously defined groups of units (if we have a pooled cross-section).
  
2. Suppose we have a panel of  $N$  units across  $T > 1$  periods and estimate  $\beta$  in the two-way fixed effects (TWFE) model  $Y_{it} = \alpha_i + \theta_t + \beta \cdot X_{it} + U_{it}$ . Why do multiple observations of units over time deal with all time-invariant omitted variables in the error term?
  
3. Discuss the assumptions needed to consistently estimate  $\beta$  in the two-way fixed effects (TWFE) model  $Y_{it} = \alpha_i + \theta_t + \beta \cdot X_{it} + U_{it}$  when  $T = 2$  via:
  - OLS on first-differenced variables  $\Delta Y_{it} = \Delta \theta_t + \beta \cdot \Delta X_{it} + \Delta U_{it}$
  - OLS on time-demeaned variables  $[Y_{it} - \bar{Y}_i] = \beta \cdot [X_{it} - \bar{X}_i] + [U_{it} - \bar{U}_i]$
  - OLS with unit/period dummy variables  $Y_{it} = \alpha_i + \theta_t + \beta \cdot X_{it} + U_{it}$

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<sup>1</sup> We focus on contemporaneous specifications and ignore more complicated patterns with lagged effects.

4. Compare the probability limit of OLS and IV when the respective exogeneity assumption  $\text{Cov}(X, U) = 0$  and  $\text{Cov}(Z, U) = 0$  may fail to be true and the IV relevance assumption  $\text{Cov}(Z, X) \neq 0$  is true, but the instrument is weak (so relevance is barely true).
  
5. Suppose we have access to an instrument  $Z$  that is relevant, exogenous, and excluded for treatment  $X$  in the linear model  $Y = \beta_0 + \beta_1 X + U$ . Using *Stata*, verify that the slope  $\beta_1$  in this model can be estimated in three equivalent ways:
  - Ratio of instrument covariance with outcome  $\widehat{\text{Cov}}(Z, Y)$  and treatment  $\widehat{\text{Cov}}(Z, X)$
  - Ratio of reduced form ( $Y = \delta_0 + \delta_1 Z + V$ ) and first stage ( $X = \pi_0 + \pi_1 Z + W$ ) slope coefficients estimated via ordinary least squares  $\frac{\hat{\delta}_1^{\text{OLS}}}{\hat{\pi}_1^{\text{OLS}}}$
  - Slope of alternative outcome model ( $Y = \tilde{\beta}_0 + \tilde{\beta}_1 \cdot \hat{X} + \tilde{U}$ ) where we use predicted treatment  $\hat{X} := \hat{\pi}_0 + \hat{\pi}_1 Z$  from first stage regression